

## 16.10: Discrete-Time Reliability Chains

### The Success-Runs Chain

Suppose that we have a sequence of *trials*, each of which results in either *success* or *failure*. Our basic assumption is that if there have been  $x \in \mathbb{N}$  consecutive successes, then the probability of success on the next trial is  $p(x)$ , independently of the past, where  $p : \mathbb{N} \rightarrow (0, 1)$ . Whenever there is a failure, we start over, independently, with a new sequence of trials. Appropriately enough,  $p$  is called the *success function*. Let  $X_n$  denote the length of the run of successes after  $n$  trials.

$\mathbf{X} = (X_0, X_1, X_2, \dots)$  is a discrete-time Markov chain with state space  $\mathbb{N}$  and transition probability matrix  $P$  given by

$$P(x, x+1) = p(x), \quad P(x, 0) = 1 - p(x); \quad x \in \mathbb{N} \quad (16.10.1)$$

The Markov chain  $\mathbf{X}$  is called the *success-runs chain*.

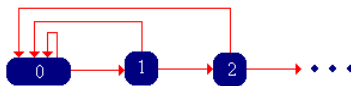


Figure 16.10.1: State graph of the success-runs chain

Now let  $T$  denote the trial number of the first failure, starting with a fresh sequence of trials. Note that in the context of the success-runs chain  $\mathbf{X}$ ,  $T = \tau_0$ , the first return time to state 0, starting in 0. Note that  $T$  takes values in  $\mathbb{N}_+ \cup \{\infty\}$ , since presumably, it is possible that no failure occurs. Let  $r(n) = \mathbb{P}(T > n)$  for  $n \in \mathbb{N}$ , the probability of at least  $n$  consecutive successes, starting with a fresh set of trials. Let  $f(n) = \mathbb{P}(T = n+1)$  for  $n \in \mathbb{N}$ , the probability of exactly  $n$  consecutive successes, starting with a fresh set of trials.

The functions  $p$ ,  $r$ , and  $f$  are related as follows:

1.  $p(x) = r(x+1)/r(x)$  for  $x \in \mathbb{N}$
2.  $r(n) = \prod_{x=0}^{n-1} p(x)$  for  $n \in \mathbb{N}$
3.  $f(n) = [1 - p(n)] \prod_{x=0}^{n-1} p(x)$  for  $n \in \mathbb{N}$
4.  $r(n) = 1 - \sum_{x=0}^{n-1} f(x)$  for  $n \in \mathbb{N}$
5.  $f(n) = r(n) - r(n+1)$  for  $n \in \mathbb{N}$

Thus, the functions  $p$ ,  $r$ , and  $f$  give equivalent information. If we know one of the functions, we can construct the other two, and hence any of the functions can be used to define the success-runs chain. The function  $r$  is the reliability function associated with  $T$ .

The function  $r$  is characterized by the following properties:

1.  $r$  is positive.
2.  $r(0) = 1$
3.  $r$  is strictly decreasing.

The function  $f$  is characterized by the following properties:

1.  $f$  is positive.
2.  $\sum_{x=0}^{\infty} f(x) \leq 1$

Essentially,  $f$  is the probability density function of  $T-1$ , except that it may be *defective* in the sense that the sum of its values may be less than 1. The leftover probability, of course, is the probability that  $T = \infty$ . This is the critical consideration in the classification of the success-runs chain, which we will consider shortly.

Verify that each of the following functions has the appropriate properties, and then find the other two functions:

1.  $p$  is a constant in  $(0, 1)$ .
2.  $r(n) = 1/(n+1)$  for  $n \in \mathbb{N}$ .
3.  $r(n) = (n+1)/(2n+1)$  for  $n \in \mathbb{N}$ .

4.  $p(x) = 1/(x+2)$  for  $x \in \mathbb{N}$ .

Answer

1.  $p(x) = p$  for  $x \in \mathbb{N}$ .  $r(n) = p^n$  for  $n \in \mathbb{N}$ .  $f(n) = (1-p)p^n$  for  $n \in \mathbb{N}$ .
2.  $p(x) = \frac{x+1}{x+2}$  for  $x \in \mathbb{N}$ .  $r(n) = \frac{1}{n+1}$  for  $n \in \mathbb{N}$ .  $f(n) = \frac{1}{n+1} - \frac{1}{n}$  for  $n \in \mathbb{N}$ .
3.  $p(x) = \frac{(x+2)(2x+1)}{(x+1)(2x+3)}$  for  $x \in \mathbb{N}$ .  $r(n) = \frac{n+1}{2n+1}$  for  $n \in \mathbb{N}$ .  $f(n) = \frac{n+1}{2n+1} - \frac{n+2}{2n+3}$  for  $n \in \mathbb{N}$ .
4.  $p(x) = \frac{1}{x+2}$  for  $x \in \mathbb{N}$ .  $r(n) = \frac{1}{(n+1)!}$  for  $n \in \mathbb{N}$ .  $f(n) = \frac{1}{(n+1)!} - \frac{1}{(n+2)!}$  for  $n \in \mathbb{N}$ .

In part (a), note that the trials are Bernoulli trials. We have an app for this case.

The success-runs app is a simulation of the success-runs chain based on Bernoulli trials. Run the simulation 1000 times for various values of  $p$  and various initial states, and note the general behavior of the chain.

The success-runs chain is irreducible and aperiodic.

Proof

The chain is irreducible, since 0 leads to every other state, and every state leads back to 0. The chain is aperiodic since  $P(0, 0) > 0$ .

Recall that  $T$  has the same distribution as  $\tau_0$ , the first return time to 0 starting at state 0. Thus, the classification of the chain as recurrent or transient depends on  $\alpha = \mathbb{P}(T = \infty)$ . Specifically, the success-runs chain is transient if  $\alpha > 0$  and recurrent if  $\alpha = 0$ . Thus, we see that the chain is recurrent if and only if a failure is sure to occur. We can compute the parameter  $\alpha$  in terms of each of the three functions that define the chain.

In terms of  $p$ ,  $r$ , and  $f$ ,

$$\alpha = \prod_{x=0}^{\infty} p(x) = \lim_{n \rightarrow \infty} r(n) = 1 - \sum_{x=0}^{\infty} f(x) \quad (16.10.2)$$

Compute  $\alpha$  and determine whether the success-runs chain  $\mathbf{X}$  is transient or recurrent for each of the [examples above](#).

Answer

1.  $\alpha = 0$ , recurrent.
2.  $\alpha = 0$ , recurrent.
3.  $\alpha = \frac{1}{2}$ , transient.
4.  $\alpha = 0$ , recurrent.

Run the simulation of the success-runs chain 1000 times for various values of  $p$ , starting in state 0. Note the return times to state 0.

Let  $\mu = \mathbb{E}(T)$ , the expected trial number of the first failure, starting with a fresh sequence of trials.

$\mu$  is related to  $\alpha$ ,  $f$ , and  $r$  as follows:

1. If  $\alpha > 0$  then  $\mu = \infty$
2. If  $\alpha = 0$  then  $\mu = 1 + \sum_{n=0}^{\infty} n f(n)$
3.  $\mu = \sum_{n=0}^{\infty} r(n)$

Proof

1. If  $\alpha = \mathbb{P}(T = \infty) > 0$  then  $\mu = \mathbb{E}(T) = \infty$ .
2. If  $\alpha = 0$ , so that  $T$  takes values in  $\mathbb{N}_+$ , then  $f$  is the PDF of  $T-1$ , so  $\mu = 1 + \mathbb{E}(T-1)$ .
3. This is a basic result from the general theory of expected value:  $\mathbb{E}(T) = \sum_{n=0}^{\infty} \mathbb{P}(T > n)$ .

The success-runs chain  $\mathbf{X}$  is positive recurrent if and only if  $\mu < \infty$ .

### Proof

Since  $T$  is the return time to 0, starting at 0, and since the chain is irreducible, it follows from the general theory that the chain is positive recurrent if and only if  $\mu = \mathbb{E}(T) < \infty$ .

If  $\mathbf{X}$  is recurrent, then  $r$  is invariant for  $\mathbf{X}$ . In the positive recurrent case, when  $\mu < \infty$ , the invariant distribution has probability density function  $g$  given by

$$g(x) = \frac{r(x)}{\mu}, \quad x \in \mathbb{N} \quad (16.10.3)$$

### Proof

If  $y \in \mathbb{N}_+$  then from the result [above](#),

$$(rP)(y) = \sum_{x=0}^{\infty} r(x)P(x, y) = r(y-1)p(y-1) = r(y) \quad (16.10.4)$$

For  $y = 0$ , using the result [above](#) again,

$$(rP)(0) = \sum_{x=0}^{\infty} r(x)P(x, 0) = \sum_{x=0}^{\infty} r(x)[1 - p(x)] = \sum_{x=0}^{\infty} [r(x) - r(x)p(x)] = \sum_{x=0}^{\infty} [r(x) - r(x+1)] \quad (16.10.5)$$

If the chain is recurrent,  $r(n) \rightarrow 0$  as  $n \rightarrow \infty$  so the last sum collapses to  $r(0) = 1$ . Recall that  $\mu = \sum_{n=0}^{\infty} r(n)$ . Hence if  $\mu < \infty$ , so that the chain is positive recurrent, the function  $g$  (which is just  $r$  normalized) is the invariant PDF.

When  $\mathbf{X}$  is recurrent, we know from the general theory that every other nonnegative left invariant function is a nonnegative multiple of  $r$

Determine whether the success-runs chain  $\mathbf{X}$  is transient, null recurrent, or positive recurrent for each of the [examples above](#). If the chain is positive recurrent, find the invariant probability density function.

Answer

1.  $\mu = \frac{1}{1-p}$ , positive recurrent.  $g(x) = (1-p)p^x$  for  $x \in \mathbb{N}$ .
2.  $\alpha = 0$ ,  $\mu = \infty$ , null recurrent.
3.  $\alpha = \frac{1}{2}$ , transient.
4.  $\mu = e - 1$ , positive recurrent.  $g(x) = \frac{1}{(e-1)(x+1)!}$  for  $x \in \mathbb{N}$ .

From (a), the success-runs chain corresponding to Bernoulli trials with success probability  $p \in (0, 1)$  has the geometric distribution on  $\mathbb{N}$ , with parameter  $1 - p$ , as the invariant distribution.

Run the simulation of the success-runs chain 1000 times for various values of  $p$  and various initial states. Compare the empirical distribution to the invariant distribution.

## The Remaining Life Chain

Consider a device whose (discrete) time to failure  $U$  takes values in  $\mathbb{N}$ , with probability density function  $f$ . We assume that  $f(n) > 0$  for  $n \in \mathbb{N}$ . When the device fails, it is immediately (and independently) replaced by an identical device. For  $n \in \mathbb{N}$ , let  $Y_n$  denote the time to failure of the device that is in service at time  $n$ .

$\mathbf{Y} = (Y_0, Y_1, Y_2, \dots)$  is a discrete-time Markov chain with state space  $\mathbb{N}$  and transition probability matrix  $Q$  given by

$$Q(0, x) = f(x), \quad Q(x+1, x) = 1; \quad x \in \mathbb{N} \quad (16.10.6)$$

The Markov chain  $\mathbf{Y}$  is called the *remaining life chain* with lifetime probability density function  $f$ , and has the state graph below.

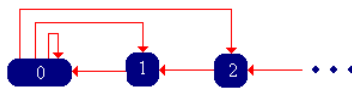


Figure 16.10.2: State graph of the remaining life chain

We have an app for the remaining life chain whose lifetime distribution is the geometric distribution on  $\mathbb{N}$ , with parameter  $1 - p \in (0, 1)$ .

Run the simulation of the remaining-life chain 1000 times for various values of  $p$  and various initial states. Note the general behavior of the chain.

If  $U$  denotes the lifetime of a device, as before, note that  $T = 1 + U$  is the return time to 0 for the chain  $\mathbf{Y}$ , starting at 0.

$\mathbf{Y}$  is irreducible, aperiodic, and recurrent.

Proof

From the assumptions on  $f$ , state 0 leads to every other state (including 0), and every positive state leads (deterministically) to 0. Thus the chain is irreducible and aperiodic. By assumption,  $\mathbb{P}(U \in \mathbb{N}) = 1$  so  $\mathbb{P}(T < \infty) = 1$  and hence the chain is recurrent.

Now let  $r(n) = \mathbb{P}(U \geq n) = \mathbb{P}(T > n)$  for  $n \in \mathbb{N}$  and let  $\mu = \mathbb{E}(T) = 1 + \mathbb{E}(U)$ . Note that  $r(n) = \sum_{x=n}^{\infty} f(x)$  and  $\mu = 1 + \sum_{x=0}^{\infty} f(x) = \sum_{n=0}^{\infty} r(n)$ .

The success-runs chain  $\mathbf{X}$  is positive recurrent if and only if  $\mu < \infty$ , in which case the invariant distribution has probability density function  $g$  given by

$$g(x) = \frac{r(x)}{\mu}, \quad x \in \mathbb{N} \quad (16.10.7)$$

Proof

Since the chain is irreducible, it is positive recurrent if and only if  $\mu = E(T) < \infty$ . The function  $r$  is invariant for  $Q$ : for  $y \in \mathbb{N}$

$$\begin{aligned} (rQ)(y) &= \sum_{x \in \mathbb{N}} r(x)Q(x, y) = r(0)Q(0, y) + r(y+1)Q(y+1, y) \\ &= f(y) + r(y+1) = r(y) \end{aligned}$$

In the positive recurrent case,  $\mu$  is the normalizing constant for  $r$ , so  $g$  is the invariant PDF.

Suppose that  $\mathbf{Y}$  is the remaining life chain whose lifetime distribution is the geometric distribution on  $\mathbb{N}$  with parameter  $1 - p \in (0, 1)$ . Then this distribution is also the invariant distribution.

Proof

By assumption,  $f(x) = (1 - p)p^x$  for  $x \in \mathbb{N}$ , and the mean of this distribution is  $p/(1 - p)$ . Hence  $\mu = 1 + p/(1 - p) = 1/(1 - p)$ , and  $r(x) = \sum_{y=x}^{\infty} f(y) = p^x$  for  $x \in \mathbb{N}$ . Hence  $g = f$ .

Run the simulation of the success-runs chain 1000 times for various values of  $p$  and various initial states. Compare the empirical distribution to the invariant distribution.

## Time Reversal

You probably have already noticed similarities, in notation and results, between the success-runs chain and the remaining-life chain. There are deeper connections.

Suppose that  $f$  is a probability density function on  $\mathbb{N}$  with  $f(n) > 0$  for  $n \in \mathbb{N}$ . Let  $\mathbf{X}$  be the success-runs chain associated with  $f$  and  $\mathbf{Y}$  the remaining life chain associated with  $f$ . Then  $\mathbf{X}$  and  $\mathbf{Y}$  are time reversals of each other.

Proof

Under the assumptions on  $f$ , both chains are recurrent and irreducible. Hence it suffices to show that

$$r(x)P(x, y) = r(y)Q(y, x), \quad x, y \in \mathbb{N} \quad (16.10.8)$$

It will then follow that the chains are time reversals of each other, and that  $r$  is a common invariant function (unique up to multiplication by positive constants). In the case that  $\mu = \sum_{n=0}^{\infty} r(n) < \infty$ , the function  $g = r/\mu$  is the common invariant PDF. There are only two cases to consider. With  $y = 0$ , we have  $r(x)P(x, 0) = r(x)[1 - p(x)]$  and  $r(0)Q(y, 0) = f(x)$ . But  $r(x)[1 - p(x)] = f(x)$  by the result [above](#). When  $x \in \mathbb{N}$  and  $y = x + 1$ , we have  $r(x)P(x, x + 1) = r(x)p(x)$  and  $r(x + 1)Q(x + 1, x) = r(x + 1)$ . But  $r(x)p(x) = r(x + 1)$  by the result [above](#).

In the context of reliability, it is also easy to see that the chains are time reversals of each other. Consider again a device whose random lifetime takes values in  $\mathbb{N}$ , with the device immediately replaced by an identical device upon failure. For  $n \in \mathbb{N}$ , we can think of  $X_n$  as the age of the device in service at time  $n$  and  $Y_n$  as the time remaining until failure for that device.

Run the simulation of the success-runs chain 1000 times for various values of  $p$ , starting in state 0. This is the time reversal of the simulation in the next exercise

Run the simulation of the remaining-life chain 1000 times for various values of  $p$ , starting in state 0. This is the time reversal of the simulation in the previous exercise.

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