

13.4: Problems on Transform Methods

Exercise 13.4.1

Calculate directly the generating function $g_X(s)$ for the geometric (p) distribution.

Answer

$$g_X(s) = E[s^2] = \sum_{k=0}^{\infty} p_k s^k = p \sum_{k=0}^{\infty} q^k s^k = \frac{p}{1-qs} \quad (\text{geometric series})$$

Exercise 13.4.2

Calculate directly the generating function $g_X(s)$ for the Poisson (μ) distribution.

Answer

$$g_X(s) = E[s^X] = \sum_{k=0}^{\infty} p_k s^k = e^{-\mu} \sum_{k=0}^{\infty} \frac{\mu^k s^k}{k!} = e^{-\mu} e^{\mu s} = e^{\mu(s-1)}$$

Exercise 13.4.3

A projection bulb has life (in hours) represented by $X \sim \text{exponential}(1/50)$. The unit will be replaced immediately upon failure or at 60 hours, whichever comes first. Determine the moment generating function for the time Y to replacement.

Answer

$$\begin{aligned} Y &= I_{[0,a]}(X)X + I_{(a,\infty)}(X)a \quad \backslash (e^{\wedge}\{sY\} = I_{-}\{[0, a)\} (X) e^{\wedge}\{sX\} + I_{-}\{(a, \infty) (X) e^{\wedge}\{as\}\}) \\ M_Y(s) &= \int_0^a e^{st} \lambda e^{-\lambda t} dt + s^a \int_a^{\infty} \lambda e^{-\lambda t} dt \\ &= \frac{\lambda}{\lambda-s} [1 - e^{-(\lambda-s)a}] + e^{-(\lambda-s)a} \end{aligned}$$

Exercise 13.4.4

Simple random variable X has distribution

$$X = [-3 -2 0 1 4] \quad P_X = [0.15 \ 0.20 \ 0.30 \ 0.25 \ 0.10]$$

- Determine the moment generating function for X
- Show by direct calculation the $M'_X(0) = E[X]$ and $M''_X(0) = E[X^2]$.

Answer

$$\begin{aligned} M_X(s) &= 0.15e^{-3s} + 0.20e^{-2s} + 0.30 + 0.25e^s + 0.10e^{4s} \\ M'_X(s) &= -3 \cdot 0.15e^{-3s} - 2 \cdot 0.20e^{-2s} + 0 + 0.25e^s + 4 \cdot 0.10e^{4s} \\ M''_X(s) &= (-3)^2 \cdot 0.15e^{-3s} + (-2)^2 \cdot 0.20e^{-2s} + 0 + 0.25e^s + 4^2 \cdot 0.10e^{4s} \end{aligned}$$

Setting $s = 0$ and using $e^0 = 1$ give the desired results.

Exercise 13.4.5

Use the moment generating function to obtain the variances for the following distributions

EXponential (λ) Gamma (α, λ) Normal (μ, σ^2)

Answer

- Exponential:

$$M_X(s) = \frac{\lambda}{\lambda - s} \quad M'_X(s) = \frac{\lambda}{(\lambda - s)^2} \quad M''_X(s) = \frac{2\lambda}{(\lambda - s)^3}$$

$$E[X] = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda} \quad E[X^2] = \frac{2\lambda}{\lambda^3} = \frac{2}{\lambda^2} \quad \text{Var}[X] = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

b. Gamma (α, λ):

$$M_X(s) = \left(\frac{\lambda}{\lambda - s}\right)^\alpha \quad M'_X(s) = \alpha \left(\frac{\lambda}{\lambda - s}\right)^{\alpha-1} \frac{\lambda}{(\lambda - s)^2} = \alpha \left(\frac{\lambda}{\lambda - s}\right)^\alpha \frac{1}{\lambda - s}$$

$$M''_X(s) = \alpha^2 \left(\frac{\lambda}{\lambda - s}\right)^\alpha \frac{1}{\lambda - s} \frac{1}{\lambda - s} + \alpha \left(\frac{\lambda}{\lambda - s}\right)^\alpha \frac{1}{(\lambda - s)^2}$$

$$E[X] = \frac{\alpha}{\lambda} \quad E[X^2] = \frac{\alpha^2 + \alpha}{\lambda^2} \quad \text{Var}[X] = \frac{\alpha}{\lambda^2}$$

c. Normal(μ, σ):

$$M_X(s) = \exp\left(\frac{\sigma^2 s^2}{2} + \mu s\right) \quad M'_X(s) = M_X(s) \cdot (\sigma^2 s + \mu)$$

$$M''_X(s) = M_X(s) \cdot (\sigma^2 s + \mu)^2 + M_X(s) \sigma^2$$

$$E[X] = \mu \quad E[X^2] = \mu^2 + \sigma^2 \quad \text{Var}[X] = \sigma^2$$

Exercise 13.4.6

The pair $\{X, Y\}$ is iid with common moment generating function $\frac{\lambda^3}{(\lambda - s)^3}$. Determine the moment generating function for $Z = 2X - 4Y + 3$.

Answer

$$M_Z(s) = e^{3s} \left(\frac{\lambda}{\lambda - 2s}\right)^3 \left(\frac{\lambda}{\lambda + 4s}\right)^3$$

Exercise 13.4.7

The pair $\{X, Y\}$ is iid with common moment generating function $M_X(s) = (0.6 + 0.4e^s)$. Determine the moment generating function for $Z = 5X + 2Y$.

Answer

$$(M_Z(s) = (0.6 + 0.4e^{5s})(0.6 + 0.4e^{2s}))$$

Exercise 13.4.8

Use the moment generating function for the [symmetric triangular distribution](#) on $(-c, c)$ as derived in the section "Three Basic Transforms".

- Obtain an expression for the symmetric triangular distribution on (a, b) for any $a < b$.
- Use the result of part (a) to show that the sum of two independent random variables uniform on (a, b) has symmetric triangular distribution on $(2a, 2b)$.

Answer

Let $m = (a + b)/2$ and $c = (b - a)/2$. If $Y \sim$ symmetric triangular on $(-c, c)$, then $X = Y + m$ is symmetric triangular on $(m - c, m + c) = (a, b)$ and

$$M_X(s) = e^{ms} M_Y(s) = \frac{e^{cs} + e^{-cs} - 2}{c^2 s^2} e^{ms} = \frac{e^{(m+c)s} + e^{(m-c)s} - 2e^{ms}}{c^2 s^2} = \frac{e^{hs} + e^{as} - 2e^{\frac{a+b}{2}s}}{(\frac{b-a}{2})^2 s^2}$$

$$M_{X+Y}(s) = \left[\frac{e^{sb} - e^{sa}}{s(b-a)} \right]^2 = \frac{e^{s2b} + e^{s2a} - 2e^{s(b+a)}}{s^2(b-a)^2}$$

Exercise 13.4.9

Random variable X has moment generating function $\frac{p^2}{(1 - qe^s)^2}$.

- Use derivatives to determine $E[X]$ and $\text{Var}[X]$.
- Recognize the distribution from the form and compare $E[X]$ and $\text{Var}[X]$ with the result of part (a).

Answer

$$[p^2(1 - qe^s)^{-2}]' = \frac{2p^2 q e^s}{(1 - qe^s)^3} \text{ so that } E[X] = 2q/p$$

$$[p^2(1 - qe^s)^{-2}]'' = \frac{6p^2 q^2 e^{2s}}{(1 - qe^s)^4} + \frac{2p^2 q e^s}{(1 - qe^s)^3} \text{ so that } E[X^2] = \frac{6q^2}{p^2} + \frac{2q}{p}$$

$$\text{Var}[X] = \frac{2q^2}{p^2} + \frac{2q}{p} - \left(\frac{2q}{p}\right)^2 = \frac{2(q^2 + pq)}{p^2} = \frac{2q}{p^2}$$

$X \sim$ negative binomial $(2, p)$, which has $E[X] = 2q/p$ and $\text{Var}[X] = 2q/p^2$.

Exercise 13.4.10

The pair $\{X, Y\}$ is independent. $X \sim$ Poisson (4) and $Y \sim$ geometric (0, 3). Determine the generating function g_Z for $Z = 3X + 2Y$.

Answer

$$g_Z(s) = g_X(s^3)g_Y(s^2) = e^{4(s^3-1)} \cdot \frac{0.3}{1 - qs^2}$$

Exercise 13.4.11

Random variable X has moment generating function

$$M_X(s) = \frac{1}{1 - 3s} \cdot \exp(16s^2/2 + 3s)$$

By recognizing forms and using rules of combinations, determine $E[X]$ and $\text{Var}[X]$.

Answer

$$X = X_1 + X_2 \text{ with } X_1 \sim \text{exponential}(1/3) \quad X_2 \sim N(3, 16)$$

$$E[X] = 3 + 3 = 6 \quad \text{Var}[X] = 9 + 16 = 25$$

Exercise 13.4.12

Random variable X has moment generating function

$$M_X(s) = \frac{\exp(3(e^s - 1))}{1 - 5s} \cdot \exp(16s^2/2 + 3s)$$

By recognizing forms and using rules of combinations, determine $E[X]$ and $\text{Var}[X]$.

Answer

$$X = X_1 + X_2 + X_3, \text{ with } X_1 \sim \text{Poisson}(3), X_2 \sim \text{exponential}(1/5), X_3 \sim N(3, 16)$$

$$E[X] = 3 + 5 + 3 = 11 \quad \text{Var}[X] = 3 + 25 + 16 = 44$$

Exercise 13.4.13

Suppose the class $\{A, B, C\}$ of events is independent, with respective probabilities 0.3, 0.5, 0.2. Consider

$$X = -3I_A + 2I_B + 4I_C$$

- Determine the moment generating functions for and use properties of moment generating functions to determine the moment generating function for X .
- Use the moment generating function to determine the distribution for X .
- Use canonic to determine the distribution. Compare with result (b).
- Use distributions for the separate terms; determine the distribution for the sum with mgsum3. Compare with result (b).

Answer

$$M_X(s) = (0.7 + 0.3e^{-3s})(0.5 + 0.5e^{2s})(0.8 + 0.2e^{4s}) = 0.12e^{-3s} + 0.12e^{-s} + 0.28 + 0.03e^s + 0.28e^{2s} + 0.03e^{3s} + 0.07e^{4s} + 0.07e^{6s}$$

The distribution is

$$X = [-3 \ -1 \ 0 \ 1 \ 2 \ 3 \ 4 \ 6] \quad PX = [0.12 \ 0.12 \ 0.28 \ 0.03 \ 0.28 \ 0.03 \ 0.07 \ 0.07]$$

```
c = [-3 2 4 0];
P = 0.1*[3 5 2];
canonic
Enter row vector of coefficients c
Enter row vector of minterm probabilities minprob(P)
Use row matrices X and PX for calculations
Call for XDBN to view the distribution
P1 = [0.7 0.3];
P2 = [0.5 0.5];
P3 = [0.8 0.2];
X1 = [0 -3];
X2 = [0 2];
X3 = [0 4];
[x,px] = mgsum3(X1,X2,X3,P1,P2,P3);
disp([X;PX;x;px]')
```

-3.0000	0.1200	-3.0000	0.1200
-1.0000	0.1200	-1.0000	0.1200
0	0.2800	0	0.2800
1.0000	0.0300	1.0000	0.0300
2.0000	0.2800	2.0000	0.2800
3.0000	0.0300	3.0000	0.0300
4.0000	0.0700	4.0000	0.0700
6.0000	0.0700	6.0000	0.0700

Exercise 13.4.14

Suppose the pair $\{X, Y\}$ is independent, with both X and Y binomial. Use generating functions to show under what condition, if any, $X + Y$ is binomial.

Answer

Binomial iff both have same p , as shown below.

$$g_{X+Y}(s) = (q_1 + p_1 s)^n (q_2 + p_2 s)^m = (q + ps)^{n+m} \quad \text{iff } p_1 = p_2$$

Exercise 13.4.15

Suppose the pair $\{X, Y\}$ is independent, with both X and Y Poisson.

- Use generating functions to show under what condition $X + Y$ is Poisson.
- What about $X - Y$? Justify your answer.

Answer

Always Poisson, as the argument below shows.

$$g_{X+Y}(s) = e^{\mu(s-1)} e^{v(s-1)} = e^{(\mu+v)(s-1)}$$

However, $Y \sim X$ could have negative values.

Exercise 13.4.16

Suppose the pair $\{X, Y\}$ is independent, Y is nonnegative integer-valued, X is Poisson and $X + Y$ is Poisson. Use the generating functions to show that Y is Poisson.

Answer

$E[X + Y] = \mu + v$, where $v = E[Y] > 0$, $g_X(s) = e^{\mu(s-1)}$ and $g_{X+Y}(s) = g_X(s) g_Y(s) = e^{(\mu+v)(s-1)}$. Division by $g_X(s)$ gives $g_Y(s) = e^{v(s-1)}$.

Exercise 13.4.17

Suppose the pair $\{X, Y\}$ is iid, binomial (6, 0.51). By the result of [Exercise 13.4.14](#)

$X + Y$ is binomial. Use mgsum to obtain the distribution for $Z = 2X + 4Y$. Does Z have the binomial distribution? Is the result surprising? Examine the first few possible values for Z . Write the generating function for Z ; does it have the form for the binomial distribution?

Answer

```
x = 0:6;
px = ibinom(6, 0.51, x);
[Z, PZ] = mgsum(2*x, 4*x, px, px);
disp([Z(1:5); PZ(1:5)]')
```

0	0.0002	% Cannot be binomial, since odd values missing
2.0000	0.0012	
4.0000	0.0043	
6.0000	0.0118	
8.0000	0.0259	
- - - - -		

$$g_X(s) = g_Y(s) = (0.49 + 0.51s)^6 \quad g_Z(s) = (0.49 + 0.51s^2)^6 (0.49 + 0.51s^4)^6$$

Exercise 13.4.18

Suppose the pair $\{X, Y\}$ is independent, with $X \sim \text{binomial}(5, 0.33)$ and $Y \sim \text{binomial}(7, 0.47)$.

Let $G = g(X) = 3X^2 - 2X$ and $H = h(Y) = 2Y^2 + Y + 3$.

- Use the mgsum to obtain the distribution for $G + H$.
- Use icalc and csort to obtain the distribution for $G + H$ and compare with the result of part (a).

Answer

```
X = 0:5;
Y = 0:7;
PX = ibinom(5,0.33,X);
PY = ibinom(7,0.47,Y);
G = 3*X.^2 - 2*X;
H = 2*Y.^2 + Y + 3;
[Z,PZ] = mgsum(G,H,PX,PY);

icalc
Enter row matrix of X-values X
Enter row matrix of Y-values Y
Enter X probabilities PX
Enter Y probabilities PY
Use array operations on matrices X, Y, PX, PY, t, u, and P
M = 3*t.^2 - 2*t + 2*u.^2 + u + 3;
[z,pz] = csort(M,P);
e = max(abs(pz - PZ)) % Comparison of p values
e = 0
```

Exercise 13.4.19

Suppose the pair $\{X, Y\}$ is independent, with $X \sim \text{binomial}(8, 0.39)$ and $Y \sim \text{uniform on } \{-1.3, -0.5, 1.3, 2.2, 3.5\}$. Let

$U = 3X^2 - 2X + 1$ and $V = Y^3 + 2Y - 3$

- Use mgsum to obtain the distribution for $U + V$.
- Use icalc and csort to obtain the distribution for $U + V$ and compare with the result of part (a).

Answer

```
X = 0:8;
Y = [-1.3 -0.5 1.3 2.2 3.5];
PX = ibinom(8,0.39,X);
PY = (1/5)*ones(1,5);
U = 3*X.^2 - 2*X + 1;
V = Y.^3 + 2*Y - 3;
[Z,PZ] = mgsum(U,V,PX,PY);
```

```

icalc
Enter row matrix of X-values  X
Enter row matrix of Y-values  Y
Enter X probabilities  PX
Enter Y probabilities  PY
Use array operations on matrices X, Y, PX, PY, t, u, and P
M = 3*t.^2 - 2*t + 1 + u.^3 + 2*u - 3;
[z,pz] = csort(M,P);
e = max(abs(pz - PZ))
e = 0

```

Exercise 13.4.20

If X is a nonnegative integer-valued random variable, express the generating function as a power series.

- a. Show that the k th derivative at $s = 1$ is

$$g_X^{(k)}(1) = E[X(X-1)(X-2)\cdots(X-k+1)]$$

- b. Use this to show the $\text{Var}[X] = g_X''(1) + g_X'(1) - [g_X'(1)]^2$.

Answer

Since power series may be differentiated term by term

$$\begin{aligned}
 g_X^{(n)}(s) &= \sum_{k=0}^{\infty} k(k-1)\cdots(k-n+1)p_k s^{k-n} \quad \text{so that} \\
 g_X^{(n)}(1) &= \sum_{k=0}^{\infty} k(k-1)\cdots(k-n+1)p_k = E[X(X-1)\cdots(X-n+1)] \\
 \text{Var}[X] &= E[X^2] - E^2[X] = E[X(X-1)] + E[X] - E^2[X] = g_X''(1) + g_X'(1) - [g_X'(1)]^2
 \end{aligned}$$

Exercise 13.4.21

Let $M_X(\cdot)$ be the moment generating function for X .

- a. Show that $\text{Var}[X]$ is the second derivative of $e^{-s\mu} M_X(s)$ evaluated at $s = 0$.
b. Use this fact to show that $X \sim N(\mu, \sigma^2)$, then $\text{Var}[X] = \sigma^2$.

Answer

$$f(s) = e^{-s\mu} M_X(s) \quad f''(s) = e^{-s\mu} [-\mu M_X'(s) + \mu^2 M_X(s) + M_X''(s) - \mu M_X'(s)]$$

Setting $s = 0$ and using the result on moments gives

$$f''(0) = -\mu^2 + \mu^2 + E[X^2] - \mu^2 = \text{Var}[X]$$

Exercise 13.4.22

Use derivatives of $M_{M_m}(s)$ to obtain the mean and variance of the negative binomial (m, p) distribution.

Answer

To simplify writing use $f(s)$ for $M_X(s)$.

$$\begin{aligned}
 f(s) &= \frac{p^m}{(1-qe^s)^m} \quad f'(s) = \frac{mp^mqe^s}{(1-qe^s)^{m+1}} \quad f''(s) = \frac{mp^mqe^s}{(1-qe^s)^{m+1}} + \frac{m(m+1)p^mq^2e^{2s}}{(1-qe^s)^{m+2}} \\
 E[X] &= \frac{mp^mq}{(1-q)^{m+1}} = \frac{mq}{p} \quad E[X^2] = \frac{mq}{p} + \frac{m(m+1)p^mq^2}{(1-q)^{m+2}}
 \end{aligned}$$

$$\text{Var}[X] = \frac{mq}{p} + \frac{m(m+1)q^2}{p^2} - \frac{m^2q^2}{p^2} = \frac{mq}{p^2}$$

Exercise 13.4.23

Use moment generating functions to show that variances add for the sum or difference of independent random variables.

Answer

To simplify writing, set $f(s) = M_X(s)$, $g(s) = M_Y(s)$, and $h(s) = M_X(s)M_Y(s)$

$$h'(s) = f'(s)g(s) + f(s)g'(s) \quad h''(s) = f''(s)g(s) + f'(s)g'(s) + f'(s)g'(s) + f(s)g''(s)$$

Setting $s = 0$ yields

$$\begin{aligned} E[X+Y] &= E[X] + E[Y] & E[(X+Y)^2] &= E[X^2] + 2E[X]E[Y] + E[Y^2] \\ E^2[X+Y] &= E^2[X] + 2E[X]E[Y] + E^2[Y] \end{aligned}$$

Taking the difference gives $\text{Var}[X+Y] = \text{Var}[X] + \text{Var}[Y]$. A similar treatment with $g(s)$ replaced by $g(-s)$ shows $\text{Var}[X-Y] = \text{Var}[X] + \text{Var}[Y]$.

Exercise 13.4.24

The pair $\{X, Y\}$ is iid $N(3,5)$. Use the moment generating function to show that $Z = 2X - 2Y + 3$ is normal (see [Example 3](#) from "[Transform Methods](#)" for general result).

Answer

$$\begin{aligned} M_{3X}(s) &= M_X(3s) = \exp\left(\frac{9 \cdot 5s^2}{2} + 3 \cdot 3s\right) & M_{-2Y}(s) &= M_Y(-2s) = \exp\left(\frac{4 \cdot 5s^2}{2} - 2 \cdot 3s\right) \\ M_Z(s) &= e^{3s} \exp\left(\frac{(45+20)s^2}{2} + (9-6)s\right) = \exp\left(\frac{65s^2}{2} + 6s\right) \end{aligned}$$

Exercise 13.4.25

Use the central limit theorem to show that for large enough sample size (usually 20 or more), the sample average

$$A_n = \frac{1}{n} \sum_{i=1}^n X_i$$

is approximately $N(\mu, \sigma^2/n)$ for any reasonable population distribution having mean value μ and variance σ^2 .

Answer

$$E[A_n] = \frac{1}{n} \sum_{i=1}^n \mu = \mu \quad \text{Var}[A_n] = \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{\sigma^2}{n}$$

By the central limit theorem, A_n is approximately normal, with the mean and variance above.

Exercise 13.4.26

A population has standard deviation approximately three. It is desired to determine the sample size n needed to ensure that with probability 0.95 the sample average will be within 0.5 of the mean value.

- Use the Chebyshev inequality to estimate the needed sample size.
- Use the normal approximation to estimate n (see [Example 1](#) from "[Simple Random Samples and Statistics](#)").

Answer

Chevyshev inequality:

$$P\left(\frac{|A_n - \mu|}{\sigma/\sqrt{n}} \geq \frac{0.5\sqrt{n}}{3}\right) \leq \frac{3^2}{0.5^2 n} \leq 0.05 \text{ implies } n \geq 720$$

Normal approximation: Use of the table in [Example 1](#) from "[Simple Random Samples and Statistics](#)" shows

$$n \geq (3/0.5)^2 3.84 = 128$$

This page titled [13.4: Problems on Transform Methods](#) is shared under a [CC BY 3.0](#) license and was authored, remixed, and/or curated by [Paul Pfeiffer](#) via [source content](#) that was edited to the style and standards of the LibreTexts platform.